

Adaptive Modelling of Price Duration and Flash Equity Dynamics

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Flash Crash

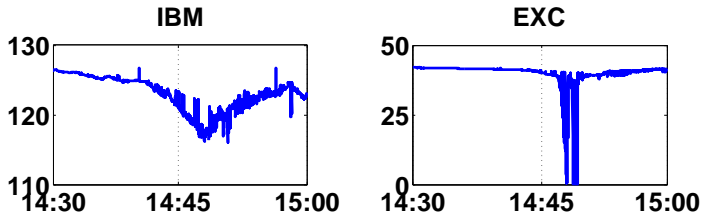


Figure 1: Stock prices of International Business Machines (IBM) and Ex-celon (EXC) during the flash crash on 20100506 at the NASDAQ stock market



Objectives

- Price Duration Dynamics
 - ▶ Local adaptive Multiplicative Error Model (MEM)
 - ▶ Balance between modelling bias and parameter variability

- Forecasting Flash Equity Dynamics
 - ▶ Estimation windows with potentially varying lengths
 - ▶ (Extreme) price movements, execution strategies



Statistics and Quantitative Finance

Statistics

- ▣ Modelling bias vs. parameter variability
- ▣ Flexible framework and predictive accuracy

Quantitative Finance Practice

- ▣ Irregularly spaced data
- ▣ Execution strategies



Example 1

Price Duration

A fund manager is trading a large number of shares of a company.

How long does it take (in seconds) that the price goes down/up by some fixed amount, say by 1%?



Example 2

Price Reversion

A fund manager observes that the price (suddenly) goes down/up.

Is a price reversion predicted? If yes, how long does the predicted recovery last (in seconds)?



Outline

1. Motivation ✓
2. Multiplicative Error Model (MEM)
3. Local Parametric Approach (LPA)
4. Price Duration and Flash Equity Dynamics
5. Conclusions




Multiplicative Error Model (MEM)

- Engle (2002), $\text{MEM}(p, q)$, \mathcal{F}_i - information set up to i

$$y_i = \mu_i \varepsilon_i, \quad E[\varepsilon_i | \mathcal{F}_{i-1}] = 1$$
$$\mu_i = \omega + \sum_{j=1}^p \alpha_j y_{i-j} + \sum_{j=1}^q \beta_j \mu_{i-j}, \quad \omega > 0, \alpha_j, \beta_j \geq 0$$

- Hautsch (2012) - comprehensive MEM literature overview
 - ▶ y_i - squared (de-means) log return: GARCH(p, q)
 - ▶ y_i - volume, bid-ask spread, duration: ACD(p, q)

Engle, Robert F. on BBI: 




Autoregressive Conditional Duration (ACD)

1. Exponential-ACD, Engle and Russel (1998) ▶ EACD

$$\varepsilon_i \sim \text{Exp}(1), \boldsymbol{\theta}_E = (\omega, \boldsymbol{\alpha}, \boldsymbol{\beta})^\top, \boldsymbol{\alpha} = (\alpha_1, \dots, \alpha_p), \boldsymbol{\beta} = (\beta_1, \dots, \beta_q)$$

2. Weibull-ACD, Engle and Russel (1998) ▶ WACD

$$\varepsilon_i \sim \mathcal{G}(s, 1), \boldsymbol{\theta}_W = (\omega, \boldsymbol{\alpha}, \boldsymbol{\beta}, s)^\top$$

Weibull, E. H. Waloddi on BBI: 



Parameter Estimation

- Consistent parameter estimation
- Data calibration with time-varying intervals
- Quasi maximum likelihood estimates (QMLEs) of θ_E and θ_W

$$\tilde{\theta}_I = \arg \max_{\theta \in \Theta} L_I(y; \theta) \quad (1)$$

- ▶ $I = [i_0 - m, i_0]$ - interval of $(m + 1)$ observations at i_0
- ▶ $L_I(\cdot)$ - log likelihood, [▶ EACD](#) and [▶ WACD](#)



Estimation Quality

- Mercurio and Spokoiny (2004), Spokoiny (2009)
- Quality of estimating *true* parameter vector θ^* by QMLE $\tilde{\theta}_I$ in terms of Kullback-Leibler divergence; $\mathcal{R}_r(\theta^*)$ - risk bound

$$E_{\theta^*} \left| L_I(\tilde{\theta}_I) - L_I(\theta^*) \right|^r \leq \mathcal{R}_r(\theta^*)$$

► Gaussian Regression

- 'Modest' risk, $r = 0.5$ (shorter intervals of homogeneity)
- 'Conservative' risk, $r = 1$ (longer intervals of homogeneity)

Kullback, Solomon and Leibler, Richard A. on BBI:



Local Parametric Approach (LPA)

- LPA, Spokoiny (1998, 2009)
 - ▶ Time series parameters can be locally approximated
 - ▶ Finding the (longest) *interval of homogeneity*
 - ▶ Balance between modelling bias and parameter variability

- Time series literature
 - ▶ GARCH(1,1) models - Čížek et al. (2009)
 - ▶ Realized volatility - Chen et al. (2010)
 - ▶ Multiplicative Error Models - Härdle et al. (2014)



Interval Selection

- $(K + 1)$ nested intervals with length $n_k = |I_k|$

$$\begin{matrix} I_0 \\ \tilde{\theta}_0 \end{matrix} \subset \begin{matrix} I_1 \\ \tilde{\theta}_1 \end{matrix} \subset \dots \subset \begin{matrix} I_k \\ \tilde{\theta}_k \end{matrix} \subset \dots \subset \begin{matrix} I_K \\ \tilde{\theta}_K \end{matrix}$$

Example: Price durations

Fix i_0 , $I_k = [i_0 - n_k, i_0]$, $n_k = [n_0 c^k]$, $c > 1$

$\{n_k\}_{k=0}^{14} = \{60 \text{ obs.}, 75 \text{ obs.}, \dots, 360 \text{ obs.}\}$, $c = 1.25$

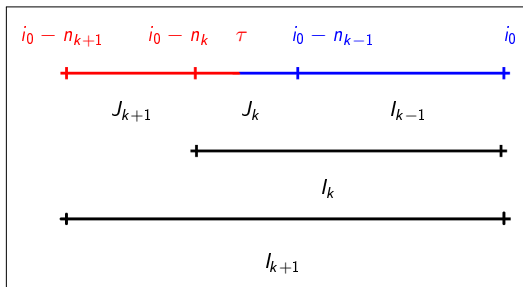


Local Change Point Detection

▶ Example

□ Fix i_0 , sequential test ($k = 1, \dots, K$)

H_0 : parameter homogeneity within I_k vs. H_1 : \exists change point within J_k



$$T_k = \sup_{\tau \in J_k} \left\{ L_{A_{k,\tau}} \left(\tilde{\theta}_{A_{k,\tau}} \right) + L_{B_{k,\tau}} \left(\tilde{\theta}_{B_{k,\tau}} \right) - L_{I_{k+1}} \left(\tilde{\theta}_{I_{k+1}} \right) \right\},$$

with $J_k = I_k \setminus I_{k-1}$, $A_{k,\tau} = [i_0 - n_{k+1}, \tau]$ and $B_{k,\tau} = (\tau, i_0]$



Critical Values, δ_k

▶ Critical Values

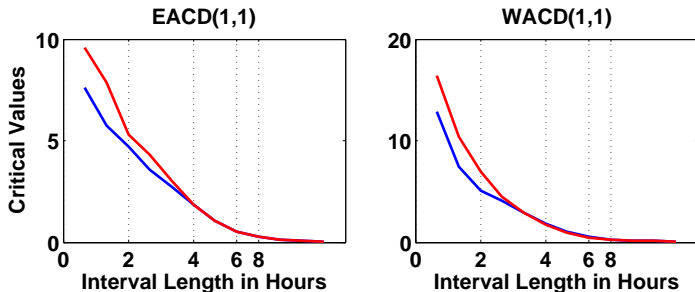


Figure 2: Critical values for low ($\tilde{\alpha} + \tilde{\beta} = 0.84$) and high ($\tilde{\alpha} + \tilde{\beta} = 0.93$) weekly persistence and 'modest' risk ($r = 0.5$) with $\rho = 0.25$, Härdle et al. (2014)



Adaptive Estimation

- Compare T_k at every step k with \mathfrak{z}_k
- Data window index of the *interval of homogeneity* - \hat{k}
- Adaptive estimate

$$\hat{\theta} = \tilde{\theta}_{\hat{k}}, \quad \hat{k} = \max_{k \leq K} \{k : T_\ell \leq \mathfrak{z}_\ell, \ell \leq k\}$$

- Note: rejecting the null at $k = 1$, $\hat{\theta}$ equals QMLE at I_0
If the algorithm goes until K , $\hat{\theta}$ equals QMLE at I_K



Data

□ NASDAQ Stock Market

- ▶ Trading days: 20100503, 20100506 (Flash crash), 20100510
- ▶ Transactions at every 25 milliseconds
- ▶ Daily trading period: 09:30-16:00
- ▶ Duration series: price moves up or down (0.95 quantile of the series)



Data

□ Companies

- ▶ Stocks without crash
Goldman Sachs (GS), Zions Bancorporation (ZION)
- ▶ Normal crash
International Business Machines (IBM), Procter&Gamble (PG)
- ▶ Penny stocks
Accenture (ACN), Excelon (EXC)



Price Duration

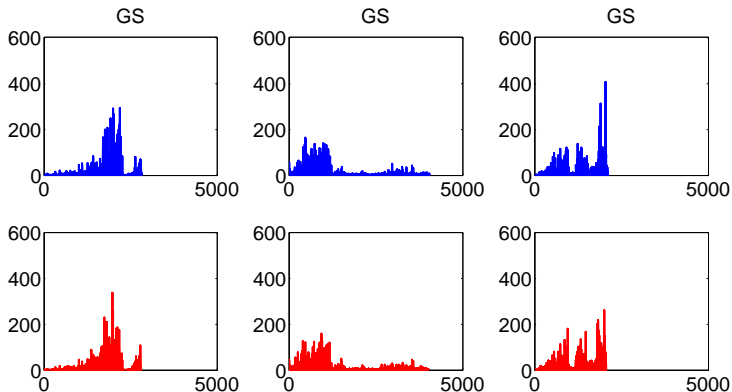


Figure 3: Price duration of GS on 20100503 (left), 20100506 (middle) and 20100510 (right). Price movement: **down** or **up**



Price Duration

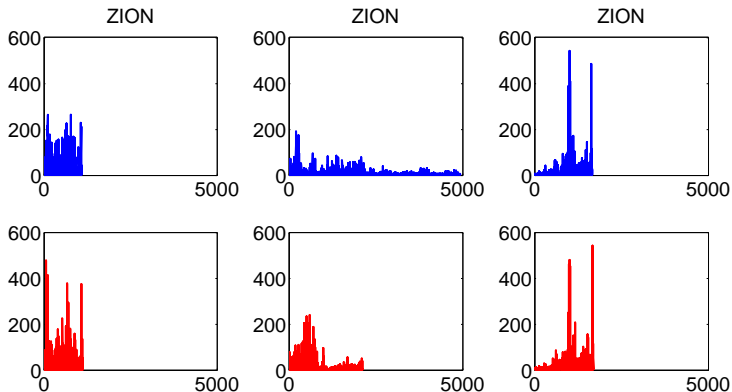


Figure 4: Price duration of ZION on 20100503 (left), 20100506 (middle) and 20100510 (right). Price movement: **down** or **up**



Price Duration

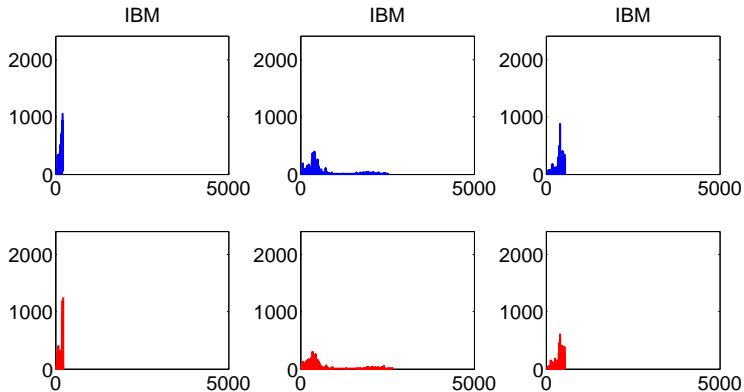


Figure 5: Price duration of IBM on 20100503 (left), 20100506 (middle) and 20100510 (right). Price movement: **down** or **up**



Price Duration

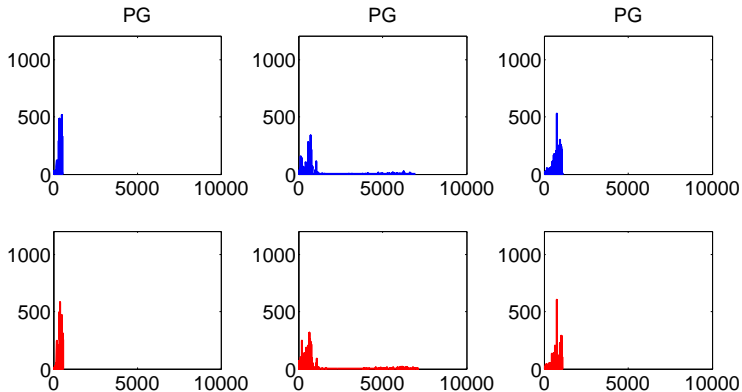


Figure 6: Price duration of PG on 20100503 (left), 20100506 (middle) and 20100510 (right). Price movement: **down** or **up**



Price Duration

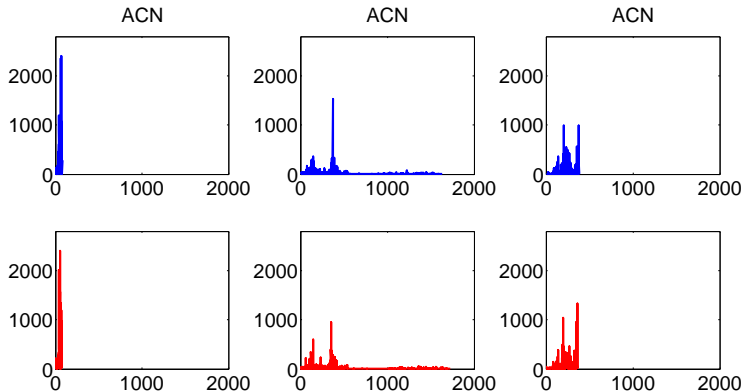


Figure 7: Price duration of ACN on 20100503 (left), 20100506 (middle) and 20100510 (right). Price movement: **down** or **up**



Price Duration

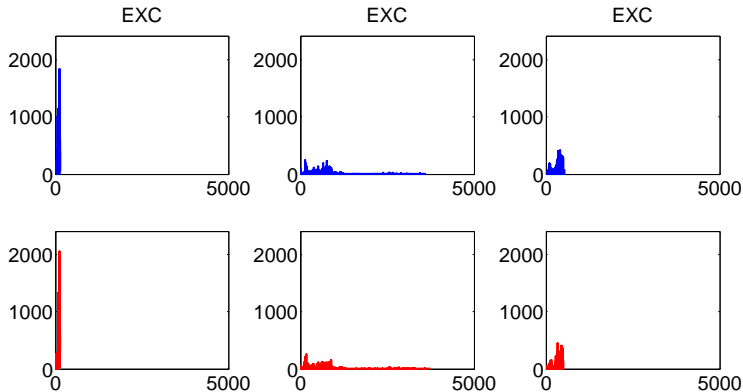


Figure 8: Price duration of EXC on 20100503 (left), 20100506 (middle) and 20100510 (right). Price movement: **down** or **up**



Conclusions

Price Duration Dynamics

- Different price duration dynamics across stocks
- Stocks without crash share stable dynamics

Forecasting Flash Equity Dynamics

- Stocks affected by the flash crash exhibit three different regimes
- Evidence for price reversion



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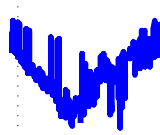
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References

-  Chen, Y. and Härdle, W. and Pigorsch, U.
Localized Realized Volatility
Journal of the American Statistical Association **105**(492):
1376–1393, 2010
-  Čížek, P., Härdle, W. and Spokoiny, V.
*Adaptive Pointwise Estimation in Time-Inhomogeneous
Conditional Heteroscedasticity Models*
Econometrics Journal **12**: 248–271, 2009
-  Engle, R. F.
New Frontiers for ARCH Models
Journal of Applied Econometrics **17**: 425–446, 2002






References

-  Engle, R. F. and Rangel, J. G.
The Spline-GARCH Model for Low-Frequency Volatility and Its Global Macroeconomic Causes
Review of Financial Studies **21**: 1187–1222, 2008
-  Engle, R. F. and Russell, J. R.
Autoregressive Conditional Duration: A New Model for Irregularly Spaced Transaction Data
Econometrica **66**(5): 1127–1162, 1998
-  Gallant, A. R.
On the bias of flexible functional forms and an essentially unbiased form
Journal of Econometrics **15**: 211–245, 1981



References

-  Härdle, W. K., Hautsch, N. and Mihoci, A.
Local Adaptive Multiplicative Error Models for High-Frequency Forecasts
Journal of Applied Econometrics, 2014
-  Hautsch, N.
Econometrics of Financial High-Frequency Data
Springer, Berlin, 2012
-  Mercurio, D. and Spokoiny, V.
Statistical inference for time-inhomogeneous volatility models
The Annals of Statistics **32**(2): 577–602, 2004



References



Spokoiny, V.

Estimation of a function with discontinuities via local polynomial fit with an adaptive window choice

The Annals of Statistics **26**(4): 1356–1378, 1998



Spokoiny, V.

Multiscale Local Change Point Detection with Applications to Value-at-Risk

The Annals of Statistics **37**(3): 1405–1436, 2009



Exponential-ACD (EACD)

▶ ACD

▶ Parameter Estimation

□ Engle and Russel (1998), $\varepsilon_i \sim \text{Exp}(1)$

$$L_I(y; \theta_E) = \sum_{i=\max(p,q)+1}^n \left(-\log \mu_i - \frac{y_i}{\mu_i} \right) \mathbf{1}\{i \in I\} \quad (2)$$

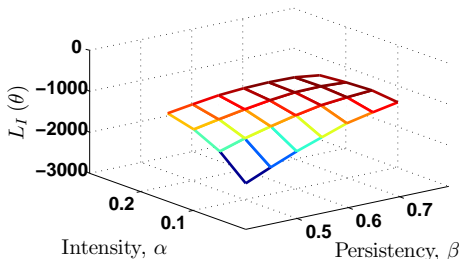


Figure 9: Log likelihood - EACD(1,1), $\theta_E^* = (0.10, 0.20, 0.65)^\top$



Weibull-ACD (WACD)

▶ ACD

▶ Parameter Estimation

□ Engle and Russel (1998), $\varepsilon_i \sim \mathcal{G}(s, 1)$

$$L_I(y; \theta_W) = \sum_{i \in I} \left[\log \frac{s}{y_i} + s \log \frac{\Gamma(1 + 1/s) y_i}{\mu_i} - \left\{ \frac{\Gamma(1 + 1/s) y_i}{\mu_i} \right\}^s \right] \mathbf{1}_{\{i \in I\}} \quad (3)$$

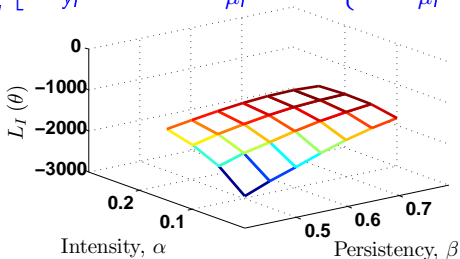


Figure 10: Log likelihood - WACD(1,1), $\theta_W^* = (0.10, 0.20, 0.65, 0.85)^\top$



Gaussian Regression

► Estimation Quality

$Y_i = f(X_i) + \varepsilon_i, i = 1, \dots, n$, weights $W = \{w_i\}_{i=1}^n$

$L(W, \theta) = \sum_{i=1}^n \ell\{Y_i, f_\theta(X_i)\} w_i$, log-density $\ell(\cdot)$, $\tilde{\theta} = \arg \max_{\theta \in \Theta} L(W, \theta)$

1. Local constant, $f(X_i) \approx \theta^*$, $\varepsilon_i \sim N(0, \sigma^2)$

$$E_{\theta^*} \left| L(W, \tilde{\theta}) - L(W, \theta^*) \right|^r \leq E |\xi|^{2r}, \quad \xi \sim N(0, 1)$$

2. Local linear, $f(X_i) \approx \theta^{*\top} \Psi_i$, $\varepsilon_i \sim N(0, \sigma^2)$, basis functions $\Psi = \{\psi_1(X_1), \dots, \psi_p(X_p)\}$, multivariate ξ

$$E_{\theta^*} \left| L(W, \tilde{\theta}) - L(W, \theta^*) \right|^r \leq E |\xi|^{2r}, \quad \xi \sim N(0, \mathcal{I}_p)$$



Local Change Point Detection ▶ LCP

Example: Price duration

- Scheme with $(K + 1) = 14$ intervals and fix i_0
- Assume $I_0 = 60$ obs. is homogeneous
- H_0 : parameter homogeneity within $I_1 = 75$ min.
 - ▶ Define $J_1 = I_1 \setminus I_0$ - observations from y_{i_0-75} up to y_{i_0-60}
 - ▶ For each $\tau \in J_1$ fit log likelihoods over $A_{1,\tau}$, $B_{1,\tau}$ and I_2
 - ▶ Find the largest likelihood ratio - T_{I_1, J_1}



Critical Values, \mathfrak{z}_k ▶ Critical Values

- Simulate \mathfrak{z}_k - homogeneity of the interval sequence l_0, \dots, l_k
- 'Propagation' condition (under H_0)

$$E_{\theta^*} \left| L_{l_k}(\tilde{\theta}_k) - L_{l_k}(\hat{\theta}_k) \right|^r \leq \rho_k \mathcal{R}_r(\theta^*), \quad k = 1, \dots, K \quad (4)$$

$\rho_k = \rho k / K$ for given significance level ρ , ▶ $\hat{\theta}_k$ - adaptive estimate

- Check \mathfrak{z}_k for (nine) different θ^* ▶ Parameter Dynamics - Quartiles
 - ▶ EACD and WACD, $K \in \{8, 13\}$, $r \in \{0.5, 1\}$, $\rho \in \{0.25, 0.50\}$
 - ▶ Findings: \mathfrak{z}_k are virtually invariable w.r.t. θ^* given a scenario
Largest differences at first two or three steps



Parameter Dynamics

Estimation window	EACD(1, 1)			WACD(1, 1)		
	Q25	Q50	Q75	Q25	Q50	Q75
1 week	0.85	0.89	0.93	0.82	0.88	0.92
2 days	0.77	0.86	0.92	0.74	0.84	0.91
1 day	0.68	0.82	0.90	0.63	0.79	0.89
3 hours	0.54	0.75	0.88	0.50	0.72	0.87
2 hours	0.45	0.70	0.86	0.42	0.67	0.85
1 hour	0.33	0.58	0.80	0.31	0.57	0.80

Table 1: Quartiles of estimated persistence levels $(\tilde{\alpha} + \tilde{\beta})$, Härdle et al. (2014)



Parameter Dynamics

▶ Critical Values

Model	Low Persistence			Moderate Persistence			High Persistence		
	Q25	Q50	Q75	Q25	Q50	Q75	Q25	Q50	Q75
EACD, $\tilde{\alpha}$	0.28	0.22	0.18	0.30	0.23	0.19	0.31	0.24	0.20
EACD, $\tilde{\beta}$	0.56	0.62	0.67	0.59	0.66	0.71	0.62	0.68	0.73
WACD, $\tilde{\alpha}$	0.28	0.21	0.17	0.30	0.23	0.18	0.32	0.24	0.19
WACD, $\tilde{\beta}$	0.54	0.60	0.65	0.58	0.65	0.70	0.60	0.68	0.74

Table 2: Quartiles of estimated ratios $\tilde{\beta} / (\tilde{\alpha} + \tilde{\beta})$ (estimation windows covering 1800 observations) conditional on the persistence level: low { EACD (0.85), WACD (0.82) }, moderate { EACD (0.89), WACD (0.88) } or high { EACD (0.93), WACD (0.92) }, Härdle et al. (2014)

